

Bond Risk Premia in Consumption-based Models

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Motivation

Fed Chair Janet Yellen: large-scale asset purchases December, 2014, Press Conference

*...we're reminding the public that we continue to hold a large stock of assets, and that is tending to push down **term premiums** in longer-term yields.*

Fed Chair Ben Bernanke: decomposition March, 2006, New York

*To the extent that the decline in forward rates can be traced to a decline in the **term premium**... the effect is financially stimulative and argues for greater monetary policy restraint... However, if the behavior of long-term yields reflects current or prospective economic conditions, the implications for policy may be quite different-indeed, quite the opposite.*

Fed Chair Alan Greenspan: conundrum June, 2005, Beijing

*That improved performance has doubtless contributed to lower inflation-related **risk premiums**, and the lowering of these **premiums** is reflected in significant declines in nominal and real long-term rates. Although this explanation contributes to an understanding of the past decade, I do not believe it explains the decline of long-term interest rates over the past year despite rising short-term rates.*

Term premium: two models & two channels

- ▶ Gaussian ATSM:
 - ▶ benchmark model
 - ▶ time-varying term premia via [price of risk](#)
- ▶ Consumption-based models with recursive preferences
 - ▶ time-varying term premia via [SV](#)

Goal of this paper: reconcile the two literatures

▶ Literature

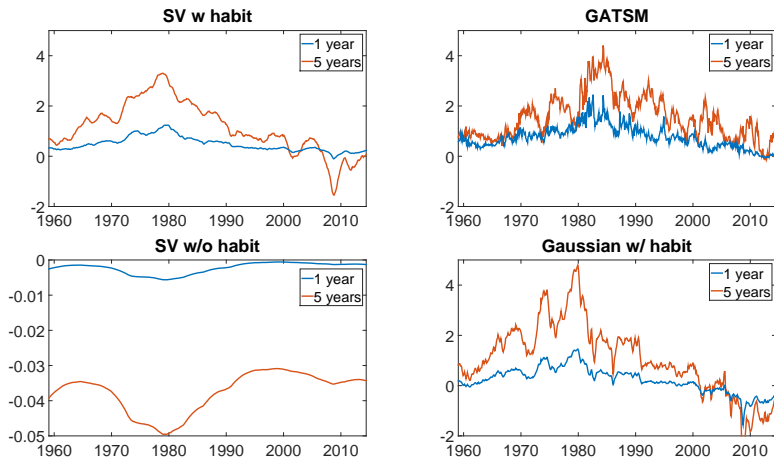
Contributions

- ▶ Introduce a new structural model with both channels
 - ▶ habit \rightarrow time-varying price of risk
 - ▶ SV \rightarrow time-varying quantity of risk
 - ▶ recursive preferences
- ▶ Our model has a reduced form of ATSM
 - ▶ inherits tractability
 - ▶ analytical bond prices
- ▶ Models with recursive preferences
 - ▶ a model solution doesn't always exist
 - ▶ we provide conditions for its existence

Empirical findings

- ▶ Our model matches empirical facts about bonds
 - ▶ realistic time variation for term premium
 - ▶ upward slope
 - ▶ mimics time series of level and slope
- ▶ Habit is the key for term premium
 - ▶ the price of expected inflation risk is the driving force
 - ▶ it comoves with the expected inflation itself
- ▶ Models with SV but not habit produce counterfactual implications for bonds
 - ▶ long run risk model
 - ▶ downward slope
 - ▶ term premia are economically insignificant, and negative

Term premia



Bottom line: habit is crucial to generate the pattern in term premia

Outline

1. Model
2. Estimation
3. Results
4. Model solution

Model

Agent's problem

$$V_t = \max_{C_t} \left[(1 - \beta) \left(\frac{C_t}{H_t} \right)^{1-\eta} + \beta \left\{ \mathbb{E}_t \left[V_{t+1}^{1-\gamma} \right] \right\}^{\frac{1-\eta}{1-\gamma}} \right]^{\frac{1}{1-\eta}}$$

s.t. $W_{t+1} = (W_t - C_t) R_{c,t+1}$

- ▶ H_t is habit
 - ▶ consumption to habit ratio enters the utility as in Abel (1999)
- ▶ β is the time discount factor
- ▶ γ measures risk aversion
- ▶ $\psi = \frac{1}{\eta}$ is the elasticity of intertemporal substitution

Stochastic discount factor

$$m_{t+1} = \vartheta \ln(\beta) + \vartheta \Delta v_{t+1} - \eta \vartheta \Delta c_{t+1} + (\vartheta - 1) r_{c,t+1}$$

- ▶ $\Delta v_{t+1} = (\eta - 1) \ln\left(\frac{H_{t+1}}{H_t}\right)$
- ▶ $\Delta c_{t+1} = \ln\left(\frac{C_{t+1}}{C_t}\right)$
- ▶ $r_{c,t+1} = \ln(R_{c,t+1})$
- ▶ $\vartheta = \frac{1-\gamma}{1-\eta}$

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- ▶ $\vartheta = \frac{1-\gamma}{1-\eta}$

$$m_{t+1}^{\$} = m_{t+1} - \pi_{t+1}$$

- ▶ π_{t+1} is inflation
- ▶ nominal variables have \$

Dynamics of the state vector

$$\Delta c_t = Z'_c g_t, \quad \pi_t = Z'_\pi g_t,$$

where

$$\begin{aligned} g_{t+1} &= \mu_g + \Phi_g g_t + \Phi_{gh} h_t + \Sigma_{gh} \varepsilon_{h,t+1} + \Sigma_{g,t} \varepsilon_{g,t+1} \\ \Sigma_{g,t} \Sigma'_{g,t} &= \Sigma_{0,g} \Sigma'_{0,g} + \sum_{i=1}^H \Sigma_{i,g} \Sigma'_{i,g} h_{it} \\ h_{t+1} &\sim \text{NCG}(\nu_h, \Phi_h, \Sigma_h) \\ \varepsilon_{h,t+1} &= h_{t+1} - \mathbb{E}_t[h_{t+1} | h_t] \end{aligned}$$

- ▶ Volatility follows non-central gamma process of Creal and Wu (JoE 2015)
- ▶ Z_c and Z_π are selection vectors
- ▶ It's a companion form, nesting long-run risk & VARMA.

Long run risk

$$\begin{aligned}
 \pi_{t+1} &= \bar{\pi}_t + \varepsilon_{\pi_1, t+1} & \varepsilon_{\pi_1, t+1} &\sim \mathbf{N}(0, h_{t, \pi_1}) \\
 \Delta c_{t+1} &= \bar{c}_t + \varepsilon_{c_1, t+1} & \varepsilon_{c_1, t+1} &\sim \mathbf{N}(0, h_{t, c_1}) \\
 \bar{\pi}_{t+1} &= \mu_\pi + \phi_\pi \bar{\pi}_t + \phi_{\pi, c} \bar{c}_t + \varepsilon_{\pi_2, t+1} & \varepsilon_{\pi_2, t+1} &\sim \mathbf{N}(0, h_{t, \pi_2}) \\
 \bar{c}_{t+1} &= \mu_c + \phi_{c, \pi} \bar{\pi}_t + \phi_c \bar{c}_t + \sigma_{c, \pi} \varepsilon_{\pi_2, t+1} + \varepsilon_{c_2, t+1} & \varepsilon_{c_2, t+1} &\sim \mathbf{N}(0, h_{t, c_2})
 \end{aligned}$$

where $g_t = (\pi_t, \Delta c_t, \bar{\pi}_t, \bar{c}_t)'$

Difference from the literature: our volatility process guarantees positivity

Habit

$$\begin{aligned}\Delta v_{t+1} &= \Lambda_1(g_t) + \Lambda_2(g_t)' \varepsilon_{g,t+1} \\ \Lambda_2(g_t) &= -\eta \Sigma_{g,t}^{-1} (\lambda_0 + \lambda_g g_t)\end{aligned}$$

- ▶ $\Lambda_2(g_t)$ is the risk sensitivity function
- ▶ $\lambda_g \neq 0 \Rightarrow$ **price of risk moves with g_t**
- ▶ **We allow inflation to be non-neutral**
 - ▶ Piazzesi and Schneider (2007) and Bansal and Shaliastovich (2013)

Relation to Wachter(2006)

If $\eta = \gamma$, $\Delta c_{t+1} = \bar{c} + \sigma_c \varepsilon_{c_1,t+1}$, $\varepsilon_{g,t+1} = \varepsilon_{c_1,t+1}$, and

$$\begin{aligned}\Lambda_{1t} &= (1 - \phi)(\bar{v} - v_t) \\ \Lambda_{2t} &= \frac{1}{\bar{H}} \sqrt{\eta + 2(v_t - \bar{v})} + \eta \sigma_c\end{aligned}$$

then the SDF becomes the same as Wachter(2006).

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The differences are

- ▶ Our model is affine
 - ▶ analytical bond prices
 - ▶ tractability
- ▶ We allow expected inflation risk to be priced
 - ▶ It turns out to be the key driving factor

Relation to preference shock

If we define $\Upsilon_t \equiv H_t^{\eta-1}$, then the objective function becomes

$$V_t = \max_{C_t} \left[(1 - \beta) \Upsilon_t C_t^{1-\eta} + \beta \left\{ E_t \left[V_{t+1}^{1-\gamma} \right] \right\}^{\frac{1-\eta}{1-\gamma}} \right]^{\frac{1}{1-\eta}},$$

where Υ_t is the time preference. The macro literature specifies

$$\Delta v_{t+1} = Z_v' g_{t+1}$$

- ▶ Z_v is a selection vector
- ▶ latent preference factor
- ▶ Albuquerque, Eichenbaum & Rebelo (2014), Schorfheide, Song & Yaron (2014)
- ▶ **no time-varying price of risk**

Model solution

Log-linearize $r_{c,t+1}$ via Campbell & Shiller (1989)

$$r_{c,t+1} \approx \kappa_0 + \kappa_1 p c_{t+1} - p c_t + \Delta c_{t+1}$$

Real pricing kernel prices consumption good

$$1 = E_t [\exp(m_{t+1} + r_{c,t+1})],$$

Guess a solution

$$p c_t = D_0 + D'_g g_t + D'_h h_t$$

Solve the fixed point problem

$$\bar{p} c = D_0 (\bar{p} c) + D'_g (\bar{p} c)' \bar{\mu}_g + D'_h (\bar{p} c)' \bar{\mu}_h$$

Plug the solution $r_{c,t+1}$ into the SDF

Sources of risk premia

Pricing kernel

$$m_{t+1}^{\$} - E_t \left[m_{t+1}^{\$} \right] = -\lambda_{g,t}^{\$, ' \Sigma_{g,t} \varepsilon_{g,t+1} - \lambda_h^{\$, ' \Sigma_{h,t} \tilde{\varepsilon}_{h,t+1}}$$

where

$$\lambda_{g,t}^{\$} = \gamma Z_c + Z_\pi \quad \leftarrow \text{power utility}$$

$$+ \kappa_1 \frac{\gamma - \eta}{1 - \eta} D_g \quad \leftarrow \text{recursive preferences}$$

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Pricing kernel

$$m_{t+1}^{\$} - E_t \left[m_{t+1}^{\$} \right] = -\lambda_{g,t}^{\$, ' \Sigma_{g,t}} \varepsilon_{g,t+1} - \lambda_h^{\$, ' \Sigma_{h,t}} \tilde{\varepsilon}_{h,t+1}$$

where

$$\begin{aligned} \lambda_{g,t}^{\$} &= \gamma Z_c + Z_\pi && \leftarrow \text{power utility} \\ &+ \kappa_1 \frac{\gamma - \eta}{1 - \eta} D_g && \leftarrow \text{recursive preferences} \\ &+ \vartheta \eta (\Sigma_{g,t} \Sigma'_{g,t})^{-1} (\lambda_0 + \lambda_g g_t) && \leftarrow \text{habit formation} \end{aligned}$$

The price of risk only varies with g_t if $\lambda_g \neq 0$.

This channel remains the same if we shut SV.

Sources of risk premia

Pricing kernel

$$m_{t+1}^{\$} - E_t \left[m_{t+1}^{\$} \right] = -\lambda_{g,t}^{\$, ' \Sigma_{g,t} \varepsilon_{g,t+1} - \lambda_h^{\$, ' \Sigma_{h,t} \tilde{\varepsilon}_{h,t+1}}$$

where

$$\begin{aligned} \lambda_h^{\$} &= \Sigma'_{gh} (\gamma Z_c + Z_{\pi}) && \leftarrow \text{power utility} \\ &+ \kappa_1 \frac{(\gamma - \eta)}{(1 - \eta)} (\Sigma'_{gh} D_g + D_h) && \leftarrow \text{recursive preference} \end{aligned}$$

Prices of volatility risks are constant, as in the literature.

Bond prices

$$P_t^{\$, (n)} = E_t \left[\exp \left(m_{t+1}^{\$} \right) P_{t+1}^{\$, (n-1)} \right]$$

yields

$$y_t^{\$, (n)} \equiv -\frac{1}{n} \ln \left(P_t^{\$, (n)} \right) = a_n^{\$} + b_{n,g}^{\$, ' } g_t + b_{n,h}^{\$, ' } h_t$$

where $b_{n,g}^{\$} = -\frac{1}{n} \bar{b}_{n,g}^{\$}$ and

$$\bar{b}_{n,g}^{\$} = \underbrace{(\Phi_g - \eta \vartheta \lambda_g)'}_{\Phi_g^{\mathbb{Q}^{\$}}} \bar{b}_{n-1,g}^{\$} + \bar{b}_{1,g}^{\$}$$

- ▶ The separation between Φ_g and $\Phi_g^{\mathbb{Q}^{\$}} \equiv \Phi_g - \eta \vartheta \lambda_g$ is the key
- ▶ Derive bond prices as in Creal & Wu (JoE 2015)
- ▶ Yields are affine functions of the state variables.
- ▶ Loadings are functions of (β, γ, ψ)

Short rate

Consumption-inflation representation

$$\begin{aligned}
 r_t^{\$} &= -\log \left(P_t^{\$, (1)} \right) \\
 &= -\ln(\beta) + \eta \mathbb{E}_t [\Delta c_{t+1}] + \mathbb{E}_t [\pi_{t+1}] \\
 &\quad - \eta \vartheta (\eta Z_c + Z_\pi)' (\lambda_0 + \lambda_g g_t) \\
 &\quad + \text{Jensen's ineq.}
 \end{aligned}$$

- ▶ Line 2: time discount, expected consumption and inflation
- ▶ **Line 3: risk adjustment**

Term premium

$$tp_t^{\$, (n)} = y_t^{\$, (n)} - \frac{1}{n} E_t \left[r_t^{\$} + r_{t+1}^{\$} + \dots + r_{t+n-1}^{\$} \right]$$

Difference between

- ▶ Buy an n -period bond
- ▶ Rolling over 1-period bond n times

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Observation equation

Stack

$$y_t^{$, (n)} = a_n^$ + b_{n,g}^{$, ' } g_t + b_{n,h}^{$, ' } h_t$$

for $n = n_1, n_2, \dots, n_N$, and allow pricing errors

$$y_t^$ = A + B_g^$ g_t + B_h^$ h_t + e_t, \quad e_t \sim \text{i.i.d.} (0, \Omega)$$

where $A^$ = (a_{n_1}^$, \dots, a_{n_N}^$)'$, $B_g^$ = (b_{g,n_1}^{$, ' }, \dots, b_{g,n_N}^{$, ' })'$, and $B_h^$ = (b_{h,n_1}^{$, ' }, \dots, b_{h,n_N}^{$, ' })'$.

Least squares

Estimate $\theta^Q = (\beta, \gamma, \psi, \Phi_g^{Q^s})$ by minimizing the pricing errors

$$\min e_t' \Omega^{-1} e_t$$

- ▶ Some macro variables g_t and h_t are latent
- ▶ We approximate $p(g_t, h_t, \theta^P | m_{1:T})$ by a Particle Gibbs sampler.

Data

Monthly data from Feb. 1959 to June 2014

Yields

- ▶ Fama-Bliss zero-coupon yields from CRSP
- ▶ maturities: 3m, 1y, 2y, 3y, 4y, 5y

Inflation + Population

- ▶ FRED database at St. Louis FRB
- ▶ CPI inflation
- ▶ Civilian population over 16

Consumption

- ▶ U.S. Bureau of Economic Analysis
- ▶ non-durables + services

Restrictions

- ▶ Four free parameters in λ_g
- ▶ $\lambda_0 = 0$
- ▶ $\Sigma_{0,g} = 0$
- ▶ Φ_h, Σ_h are diagonal
- ▶ $\Phi_{gh} = 0$ and $\Sigma_{gh} = 0$
- ▶ $\Omega = \omega^2 I$

Posterior distribution of macro factors

MCMC + particle filters \rightarrow Particle Gibbs sampler.

For $j = 1, \dots, M$

$$\begin{aligned} (g_{1:T}, h_{0:T})^{(j)} &\sim p\left(g_{1:T}, h_{0:T} \mid m_{1:T}, \theta^{\mathbb{P},(j-1)}\right) \\ \theta^{\mathbb{P},(j)} &\sim p\left(\theta^{\mathbb{P}} \mid m_{1:T}, g_{1:T}^{(j)}, h_{0:T}^{(j)}\right) \end{aligned}$$

- ▶ Draw the state variables using the particle filter, see Andrieu, Doucet, Holenstein (10).
- ▶ Use independence Metropolis-Hastings to draw the parameters $\theta^{\mathbb{P}}$.

Least squares

$$\min e_t' \Omega^{-1} e_t$$

where

$$e_t = y_t^{\$} - A^{\$}(\theta^{\text{Q}}, \hat{\theta}^{\text{P}}) + B_g^{\$}(\theta^{\text{Q}}, \hat{\theta}^{\text{P}}) \hat{g}_t + B_h^{\$}(\theta^{\text{Q}}, \hat{\theta}^{\text{P}}) \hat{h}_t$$

Structural parameters

	global		local		
Preference	ψ	1.02 (0.03)		0.70 (0.04)	
	β	0.9998 (0.0000)		1.003 (0.000)	
	γ	6.75 (2.02)		1.73 (0.16)	
Habit	$\Phi_g^{Q^s}$	0.993 (0.002)	0.018 (0.007)	0.994 (0.003)	-0.015 (0.005)
		0.000 (0.001)	0.997 (0.000)	-0.005 (0.002)	0.996 (0.000)
		λ_g	$1e^{-3} \times$		
		0.05 0.00	-0.12 0.18	-0.007 0.001	0.030 -0.023

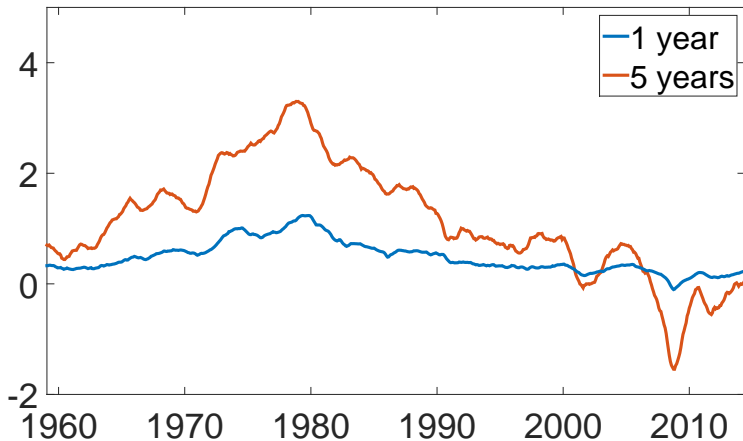
- ▶ Both global and local have similar implications for bonds.
- ▶ **Key: $\Phi_g^{Q^s}$ are persistent**
- ▶ Other structural parameters (γ, ψ, β) vary with different economic interpretations.

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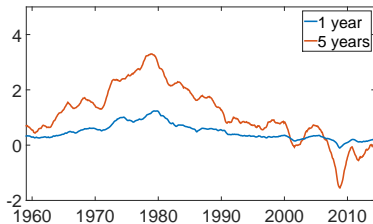
Term premia in the benchmark model

SV w habit

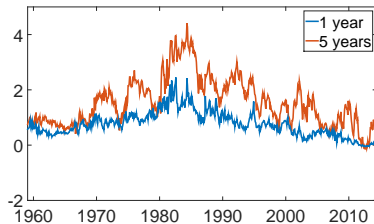


Comparison with GATSM

SV w habit

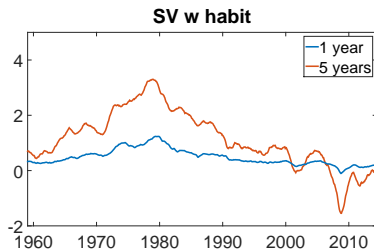
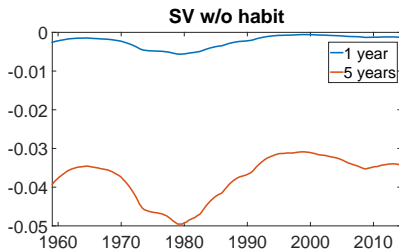


GATSM



- ▶ Left: our benchmark
- ▶ Right: GATSM

Only quantity of risk

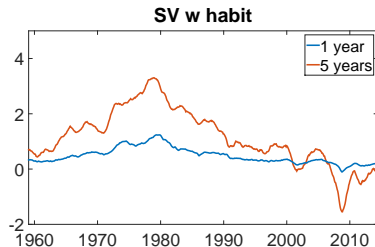
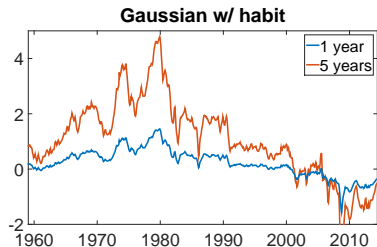


- ▶ Left: with SV, no habit (long run risk model)
- ▶ Right: our benchmark

Long run risk model produces counterfactual term premia

- ▶ economically insignificant
- ▶ negative

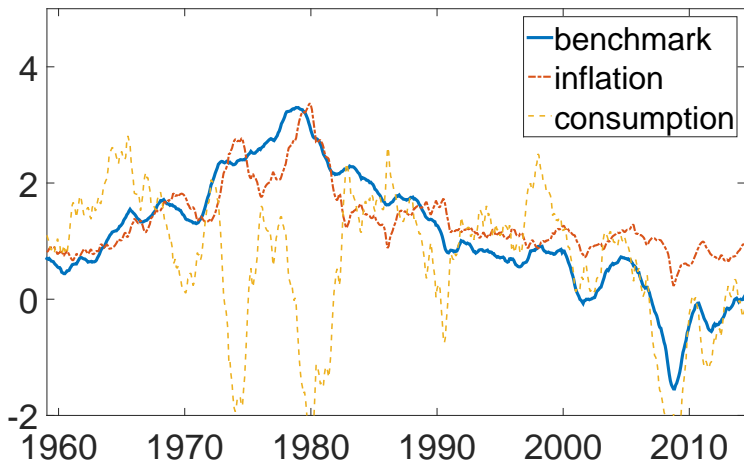
Only price of risk



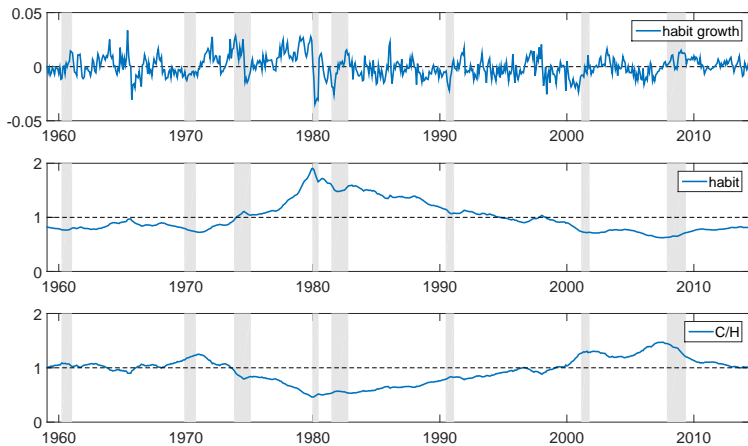
- ▶ Left: no SV, with habit
- ▶ Right: our benchmark

Bottom line: habit is crucial to generate the pattern in term premia

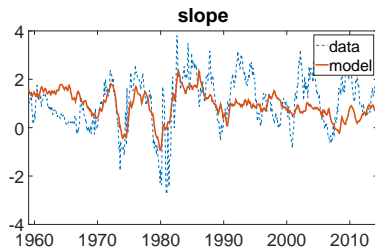
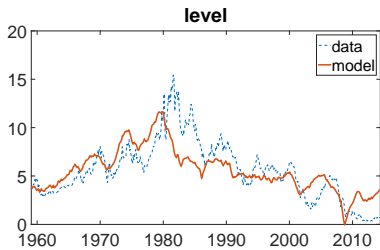
Inflation vs. consumption



Habit



Level and slope from benchmark model



- ▶ Level: average across maturities
- ▶ Slope: 5 year - 3 month

Slope

Unconditional slope has been the focus for the majority of the literature

		3	12	24	36	48	60	level	slope
data		4.94	5.33	5.54	5.72	5.88	5.98	5.57	1.04
SV w/ habit	global	4.91	5.27	5.63	5.85	5.92	5.84	5.57	0.93
	local	4.95	5.20	5.49	5.74	5.95	6.13	5.58	1.18
Gaussian w/ habit		5.08	5.25	5.47	5.69	5.89	6.09	5.58	1.01
SV w/o habit		5.64	5.63	5.61	5.59	5.57	5.56	5.60	-0.08

- ▶ SV seems to be flexible with g_t and h_t
- ▶ But there are more moments to match A , B_g , B_h
- ▶ There are only 3 free parameters (β, γ, ψ) to match all
- ▶ It's difficult to match both the average level, and slope

▶ Literature

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Model solution

Log-linearize $r_{c,t+1}$ via Campbell & Shiller (1989)

$$r_{c,t+1} \approx \kappa_0 + \kappa_1 p c_{t+1} - p c_t + \Delta c_{t+1}$$

Guess a solution

$$p c_t = D_0 + D'_g g_t + D'_h h_t$$

Solve the fixed point problem

$$\bar{p}c = D_0(\bar{p}c) + D'_g(\bar{p}c)' \bar{\mu}_g + D'_h(\bar{p}c)' \bar{\mu}_h$$

Plug the solution $r_{c,t+1}$ into the SDF

Problem: a solution to the fixed point problem does not always exist.

General case

Assumption

The parameters $\theta \in \Theta^r$ must satisfy that for any real $\bar{p}c$,

- 1. the loadings $D_h(\bar{p}c, \theta)$ are real,*
- 2. the expectation $1 = E_t[\exp(m_{t+1} + r_{c,t+1})]$ exists for $D_h(\bar{p}c, \theta)$.*

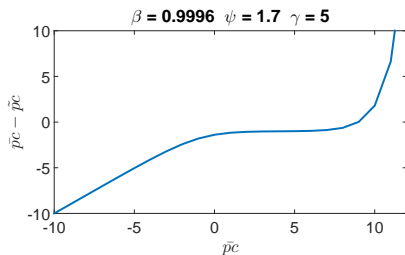
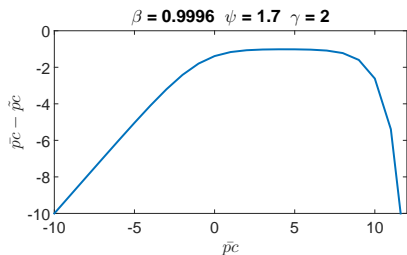
Proposition

Given Assumptions, there is a value $\bar{\beta}(\psi, \gamma, \theta^{\mathbb{P}}, \theta^{\lambda})$ such that if $\beta < \bar{\beta}$, then there exists a real solution for the fixed point problem.

Sketch of proof

Define $\tilde{\rho}c(\bar{\rho}c) = D_0(\bar{\rho}c) + D_g(\bar{\rho}c)' \bar{\mu}_g + D_h(\bar{\rho}c)' \bar{\mu}_h$

The fixed point problem has a solution if $\bar{\rho}c - \tilde{\rho}c(\bar{\rho}c) = 0$

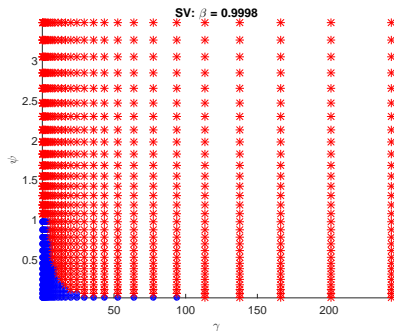
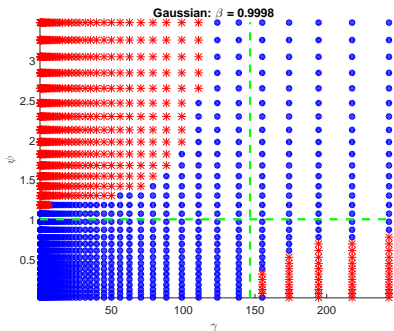


Special case: Gaussian

Corollary

1. If $Z_1^{\infty'} \mu_g^* \leq 0$ and $\beta \leq 1$, then $\frac{1-\gamma}{1-\psi} > 0$ guarantees the existence of a solution.
2. If $\beta \leq 1$, then there is a value $\bar{\gamma}(\theta^{\mathbb{P}}, \theta^{\lambda})$ such that $\frac{\bar{\gamma}-\gamma}{1-\psi} > 0$ guarantees a solution.
3. For any ψ , $\bar{\beta}$ is monotonic in γ : for $\psi > 1$, then $\frac{d\bar{\beta}}{d\gamma} > 0$; for $\psi < 1$, then $\frac{d\bar{\beta}}{d\gamma} < 0$.

Numerical illustration: part 2



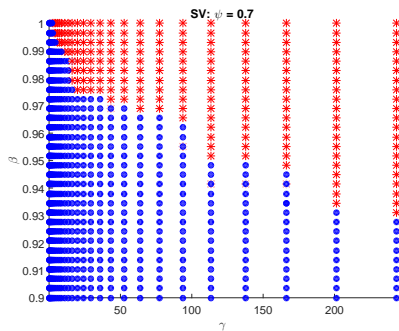
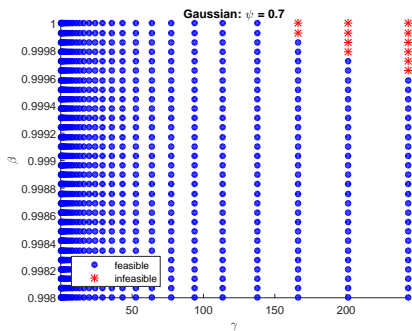
- ▶ Gaussian: $\bar{\gamma} = 146.5$
- ▶ benchmark: for $\psi = 0.97, \gamma < 4.8$
- ▶ benchmark: for $\psi = 0.52, \gamma < 6.9$

Special case: Gaussian

Corollary

1. If $Z_1^{\infty'} \mu_g^* \leq 0$ and $\beta \leq 1$, then $\frac{1-\gamma}{1-\psi} > 0$ guarantees the existence of a solution.
2. If $\beta \leq 1$, then there is a value $\bar{\gamma}(\theta^{\mathbb{P}}, \theta^{\lambda})$ such that $\frac{\bar{\gamma}-\gamma}{1-\psi} > 0$ guarantees a solution.
3. For any ψ , $\bar{\beta}$ is monotonic in γ : for $\psi > 1$, then $\frac{d\bar{\beta}}{d\gamma} > 0$; for $\psi < 1$, then $\frac{d\bar{\beta}}{d\gamma} < 0$.

Numerical illustration: part 3



- ▶ Gaussian: for $\gamma = 244, \beta < 0.9996$
- ▶ benchmark: for $\gamma = 244, \beta < 0.93$

What we have learned

- ▶ A small γ might not mark the success of a model, but simply to satisfy the constraint
- ▶ The separation of regions might cause numerical problems for estimation, frequentist or Bayesian
- ▶ SV models encounter more problems

Conclusion

- ▶ Build a new structural model with two forces for term premia
 - ▶ **habit** → time-varying prices of risk
 - ▶ SV → time-varying quantity of risk
- ▶ Empirical results:
 - ▶ Our model
 - ▶ captures realistic dynamics for risk premia
 - ▶ upward slope
 - ▶ Habit is the driving force for term premia
 - ▶ the price of expected inflation risk is the key, which comoves with the expected inflation itself
 - ▶ Models with SV but not habit produce counterfactual implications
 - ▶ downward slope
 - ▶ term premia are economically insignificant, and negative
- ▶ Provide conditions guaranteeing a solution for models with recursive preferences

Literature:

Consumption-based models

- ▶ recursive preferences: *Piazzesi Schneider (07), Le Singleton (10)*
- ▶ recursive preferences + SV: *Bansal Yaron (04), Bansal Gallant Tauchen (07), Bansal Shaliastovich (13),*
- ▶ habit formation: *Wachter (06)*
- ▶ recursive preferences + SV in ICAPM: *Campbell Giglio et al. (14)*
- ▶ recursive preferences + preference shocks: *Albuquerque Eichenbaum Rebelo (14) Schorfheide Song Yaron (14)*

DSGE models

- ▶ habit formation: *Rudebusch Swanson (08)*
- ▶ recursive preferences: *Rudebusch Swanson (08), van Binsbergen et al. (12), Dew-Becker (14)*
- ▶ solution methods: *Caldara et al. (12)*

Literature:

Term premium

- ▶ ATSM: *Duffee (02), Ang Piazzesi (03), Wright (11), Bauer Rudebusch Wu (12)*

Model solution

- ▶ *Hansen Scheinkman (12), Campbell Giglio et al. (14)*

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Comparison with estimation in the literature

$$\Phi_g^{\mathbb{Q}^s} \equiv \Phi_g - \eta^{\vartheta} \lambda_g$$

- ▶ Dynamics of macro variables: Φ_g
- ▶ Cross section of yields: $\Phi_g^{\mathbb{Q}^s}$
- ▶ Term premia: the difference between \mathbb{P} and \mathbb{Q}

In models without habit, $\Phi_g = \Phi_g^{\mathbb{Q}^s}$

- ▶ If we extract macro dynamics from macro data (*ours*), then
 - ▶ Macro factors retain their interpretation
 - ▶ Φ_g is estimated from the macro dynamics
 - ▶ and determines the slope is downward
- ▶ If we extract macro dynamics primarily from yields (*literature*), then
 - ▶ $\Phi_g^{\mathbb{Q}}$ is estimated from the cross section of yields
 - ▶ and then determines the dynamics of the factors
 - ▶ macro factors mimic level, slope and curvature of yields

Habit allows $\Phi_g \neq \Phi_g^{\mathbb{Q}}$

Macro factors and yields

Regression R^2 s of macro factors on yields

	our estimates	inversion	
		w/o p.e	w/ p.e.
expected inflation	57%	100%	98%
expeted growth	31%	100%	96%
expected inflation vol	48%	100%	36%
expected growth vol	31%	100%	72%

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